#### 36-709: Advanced Statistical Theory

Lecture 6: February 14

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Spring 2019

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### 6.1 Notes on Regression

Let  $Y_1, \dots, Y_n$  be independent response variable in  $\mathbb{R}$ . In particular,

$$Y_i = \mathsf{E}[Y_i] + \epsilon_i, \qquad \epsilon_i \in \mathrm{SG}(\sigma^2)$$

Suppose we observed *n* covariates  $(x_1, \dots, x_n) \subseteq \mathbb{R}^d$ ; fixed. Further assume that the first coordinate of each  $x_i$  be 1. Consider a following *model*:

$$Y_i = f(X_i) + \eta_i, \qquad f : \mathbb{R}^d \to \mathbb{R}, \qquad \mathsf{E}(\eta_i) = 0$$

The model specification need not be linear, however we often choose  $f(x_i) = x_i^\top \gamma$ ,  $\gamma \in \mathbb{R}^d$  for the brevity. Under the squared error loss, our goal is to solve the following problem:

$$\min_{\gamma} \sum_{i=1}^{n} \mathsf{E}(y_i - x_i^{\top} \gamma)^2$$

Let  $\beta$  be the minimizer of the problem. Assuming  $\Sigma_n = \frac{1}{n} \sum_{i=1}^n x_i x_i^{\top}$  be full rank,

$$\beta = \sum_{n=1}^{n} \frac{1}{n} \sum_{i=1}^{n} x_i \mu_i = (X^{\top} X)^{-1} X^{\top} \mu$$

where  $\mathbf{Y} = (Y_1, \dots, Y_n)^{\top}$ ,  $\mathbf{X} = (x_1, \dots, x_n)^{\top} \subseteq \mathbb{R}^{n \times d}$ ,  $\boldsymbol{\mu} = (\mu_1, \dots, \mu_n)^{\top}$  and  $\mu_i := \mathsf{E}[Y_i]$ . With empirical risk minimization, we estimate  $\beta$  with  $\hat{\beta} = (\mathbf{X}^{\top} \mathbf{X})^{-1} \mathbf{X}^{\top} \mathbf{Y}$ . Note that the excess risk  $R(\gamma)$  becomes

$$R(\gamma) = \mathsf{E}\left[\frac{1}{n}\sum_{i=1}^{n} (Y_{i} - x_{i}^{\top}\gamma)^{2} - \frac{1}{n}\sum_{i=1}^{n} (Y_{i} - x_{i}^{\top}\beta)^{2}\right]$$

See [HKZ12] for the probabilistic bound on  $R(\gamma)$  and details.

# 6.2 Concentration of *L*-Lipschitz function of Gaussian vectors

Recall from Mill's ratio,

$$X \sim \mathcal{N}(\mu, \sigma^2) \implies \mathsf{P}(|X - \mu| \ge t) \le 2 \exp\left(-\frac{t^2}{2\sigma^2}\right), \qquad t > 0$$

**Definition 6.1 (Lipschitz Condition)**  $f : \mathbb{R}^d \to \mathbb{R}$  satisfies L-Lipschitz condition with respect to Euclidean norm, if

$$\exists L \in \mathbb{R} \ s.t. \ |f(\boldsymbol{x}) - f(\boldsymbol{y})| \leq L \|\boldsymbol{x} - \boldsymbol{y}\|, \qquad \forall \boldsymbol{x}, \boldsymbol{y} \in \mathbb{R}^d$$

**Theorem 6.2** Let  $f : \mathbb{R}^d \to \mathbb{R}$  be L-Lipschitz. Suppose  $\mathbf{Z} := (Z_1, \cdots, Z_d)^\top \sim \mathrm{N}_d(\mathbf{0}, \sigma^2 I_d)$ . Then

$$\mathsf{P}\left(\left|f(\boldsymbol{Z}) - \mathsf{E}[f(\boldsymbol{Z})]\right| \ge t\right) \le 2\exp\left(-\frac{t^2}{2\sigma^2 L^2}\right), \qquad t > 0$$

That is,  $f(\mathbf{Z}) \in SG(\sigma^2 L^2)$ 

The take-home message is that the concentration of any *L*-Lipschitz function of isotropic Gaussian random vector is like a scalar Gaussian variable with variance  $\sigma^2 L^2$  independent to the dimension *d*.

*Remark.* Historically,  $\mathsf{E}(\cdot)$  was  $\mathsf{M}(\cdot)$  where  $\mathsf{M}$  denotes a median operator. See [BLM13] for detail.

*Remark.* Theorem 6.2 holds even when  $d \to \infty$ , which essentially is an isotropic Gaussian Process.

**Corollary 6.3** Let  $\mathbf{X} = (X_1, \dots, X_d)$  be i.i.d. random variables with  $X_i \sim \text{Unif}[0, 1]$ . Suppose  $f : \mathbb{R}^d \to \mathbb{R}$  be L-Lipschitz. Then,

$$\mathsf{P}\left(\left|f(\boldsymbol{X}) - \mathsf{E}[f(\boldsymbol{X})]\right| \ge t\right) \le 2\exp\left(-\frac{t^2}{2L_{\alpha}^2}\right)$$

where  $L_{\alpha}$  depends on L.

**Proof:** Use inverse CDF method.

### Maxima of Gaussians

Lemma 6.4 Let  $\mathbf{Y} := (Y_1, \cdots, Y_d)^\top \sim \mathcal{N}_d(\mathbf{0}, \Sigma), \ \Sigma \in S_d^+$  $\mathsf{P}\left(\left|\max_i Y_i - \mathsf{E}(\max_i Y_i)\right| \ge t\right) \le 2\exp\left(-\frac{t^2}{2\sigma_{\max}^2}\right), \qquad t > 0$ 

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where  $\sigma_{\max}^2 = \max_i \Sigma_{ii}$ 

**Proof:** Let  $\boldsymbol{Y} = \Sigma^{1/2} \boldsymbol{z}$  where  $\Sigma^{1/2}$  is a square root of  $\Sigma$ , i.e,  $\Sigma = \Sigma^{1/2} (\Sigma^{1/2})^{\top}$  and  $\boldsymbol{z} \sim N_d(\boldsymbol{0}, I_d)$ . *Claim:*  $f : \boldsymbol{x} \in \mathbb{R}^d \mapsto \max_i (\Sigma^{1/2} \boldsymbol{x})_i$  is *L*-Lipschitz with  $L = \max_i \sqrt{\Sigma_{ii}}$ 

Fix coordinate i and let  $\Sigma_i^{1/2}$  be the *i*th row of  $\Sigma^{1/2}$ . For all  $\boldsymbol{x}, \boldsymbol{y} \in \mathbb{R}^d$ 

$$\begin{aligned} \left| (\Sigma^{1/2} \boldsymbol{x})_i - (\Sigma^{1/2} \boldsymbol{y})_i \right| &= \left| \Sigma_i^{1/2} (\boldsymbol{x} - \boldsymbol{y}) \right| \le \sqrt{\Sigma_i^{1/2} \left( \Sigma_i^{1/2} \right)^\top} \cdot \| \boldsymbol{x} - \boldsymbol{y} \| \qquad (\because \text{Cauchy-Schwartz Ineq.}) \\ &= \sqrt{\Sigma_{ii}} \cdot \| \boldsymbol{x} - \boldsymbol{y} \| \end{aligned}$$

**Theorem 6.5 (Talagrand)** Let  $\mathbf{X} = (X_1, \dots, X_d)^{\top}$  be independent random variable with  $X_i \in [0, 1], \forall i$ . Suppose  $f : \mathbb{R}^d \to \mathbb{R}$  be L-Lipschitz with respect to Euclidean norm and f be convex function. Then,

$$\mathsf{P}\left(|f(\boldsymbol{X}) - \mathsf{E}[f(\boldsymbol{X})]| \ge t\right) \le 2\exp\left(-\frac{t^2}{2L^2}\right), \qquad t > 0$$

As in Theorem 6.2, E can be replaced to M. Theorem 6.5 is a corollary of *Convex Distance Inequality* [T95]. See [BLM13] for detail.

**Theorem 6.6 (Convex Distance Inequality)** For any subset  $A \subseteq \mathcal{X}^n$ 

$$\mathsf{P}(X \in A)\mathsf{P}\left(\sup_{\alpha \in [0,\infty)^n : \|\alpha\|=1} d_{\alpha}(X,A) \ge t\right) \le \exp(-t^2/4)$$

where  $d_{\alpha}(x, A) = \inf_{y \in A} \sum_{i: x_i \neq y_i} |\alpha_i|$  s.t.  $\alpha \in [0, \infty)^n$ ; weighted Hamming distance.

## 6.3 Prelude to the Next topic

Until now, we consider index set  $\mathcal{I}$  to be finite, i.e.,  $|\mathcal{I}| < \infty$ . However, often we need to bound either in probability or in expectation of the form  $\max_{i \in \mathcal{I}} X_i$  or  $\max_{i \in \mathcal{I}} |X_i|$  where  $|\mathcal{I}| = \infty$  under the assumption  $X_i \in SG(\cdot)$  or  $SE(\cdot)$ . One strategy is to approximate  $\max_{i \in \mathcal{I}} X_i$  with finite covers of a ball where the maximum within the ball be bounded by that within the covers.

## References

- [HKZ12] HSU, D. and KAKADE, S. M. and ZHANG, T. (2012) A tail inequality for quadratic forms of subgaussian random vectors, Electron. Commun. Prob., 17, No.52 pp. 1–6.
- [BLM13] BOUCHERON, S., LUGOSI, G., and MASSART, P. (2013) Concentration inequalities: A nonasymptotic theory of independence. Oxford university press.
  - [T95] TALAGRAND, M. (1995) Concentration of measure and isoperimetric inequalities in product spaces. Publications Mathématiques de l'Institut des Hautes Etudes Scientifiques, 81(1), 73-205.