#### 36-709: Advanced Statistical Theory

Lecture 8: February 21

Lecturer: Alessandro Rinaldo

 $Scribes:\ Shenghao\ Wu$ 

Spring 2019

Note: LaTeX template courtesy of UC Berkeley EECS dept.

**Disclaimer**: These notes have not been subjected to the usual scrutiny reserved for formal publications. They may be distributed outside this class only with the permission of the Instructor.

This lecture's notes illustrate some uses of various IATFX macros. Take a look at this and imitate.

## 8.1 Euclidean norm of sub-Gaussian random vectors

**Definition 8.1** (Sub-Gaussian random vectors)A random vector  $X \in \mathbb{R}^d$  is a sub-Gaussian random vector with parameter  $\sigma^2$  if

$$v^{T} X \in SG(\sigma^{2}), \forall v \in \mathbb{S}^{a-1}$$

where  $\mathbb{S}^{d-1} = \{x \in \mathbb{R}^d : ||x|| = 1\}$  is the d-1 unit sphere. We write  $X \in SG_d(\sigma^2)$ .

**Lemma 8.2**  $X \in \mathbb{R}^d$  is a sub-Guassian random vector with parameter  $||\Sigma||_{op}$  if  $X \sim \mathcal{N}(0, \Sigma)$ 

**Proof:** For any  $v \in \mathbb{S}^{d-1}$ ,  $v^T \Sigma v \leq ||\Sigma||_{op}$ . Take MGF:  $\mathbb{E}[e^{\lambda v^T X}] = e^{\lambda^2 v^T \Sigma v/2} \leq e^{\lambda^2 ||\Sigma||_{op}/2}$  Notice that sub-Guassian vector does not need to be a vector of independent Gaussians (but the vice is true).

We now prove the theorem from last time:

**Theorem 8.3** Let  $X \in SG_d(\sigma^2), ||X|| = \sqrt{\sum_{i=1}^d X_i^2}$ , then:

 $\mathbb{E}[||X||] \le 4\sigma\sqrt{d}$ 

Moreover, with probability at least  $1 - \delta$  for  $\delta \in (0, 1)$ :

$$||X|| \le 4\sigma\sqrt{d} + 2\sigma\sqrt{\log(\frac{1}{\delta})}$$

**Proof:** Let  $N_{\frac{1}{2}}$  be a  $\frac{1}{2}$ -minimal cover of  $B_d$  in Euclidian norm, that is:

$$\forall \theta \in B_d, \exists z = z(\theta) \in N_{\frac{1}{2}} \ s.t. \ ||\theta - z|| \le \frac{1}{2}$$

. Equivalently,  $\forall \theta \in B_d$ , we can write  $\theta = z + w$  where  $z = z(\theta) \in N_{\frac{1}{2}}$  and  $||w|| \leq \frac{1}{2}$ . Also, by the volumetric rate bounds,

$$|N_{\frac{1}{2}}| \le (1 + \frac{2}{1/2})^d = 5^d$$

Hence,

$$\max_{v \in B_d} v^T X \le \max_{z \in N_{\frac{1}{2}}} z^T X + \max_{w \in \frac{1}{2}B_d} w^T X = \max_{z \in N_{\frac{1}{2}}} z^T X + \frac{1}{2} \max_{w \in B_d} w^T X$$

Hence  $\underbrace{\max_{v \in B_d} v^T X}_{||X||} \le 2 \max_{z \in N_{\frac{1}{2}}} z^T X.$ 

In general, some argument will lead to the following bound:

$$||X|| \le \frac{1}{1-\epsilon} \max_{z \in N_{\frac{1}{2}}} z^T X \text{ for } \epsilon \in (0,1)$$

Therefore,

$$\mathbb{E}[||X||] \leq 2\mathbb{E}[\max_{z \in N_{\frac{1}{2}}} \underbrace{z^T X}_{SG(\sigma^2)}] \leq 2\sigma \sqrt{2log|N_{\frac{1}{2}}|} \leq 2\sigma \sqrt{2dlog5} \leq 4\sigma \sqrt{d}$$

The second inequality is due to the maximal inequality for sub-Gaussian random variables we have proved in class.

To prove the high probability bound, for t > 0:

$$\mathbb{P}(||X|| \ge t) \le \mathbb{P}(\max_{z \in N_{\frac{1}{2}}} z^T X \ge \frac{t}{2}) \le |N_{\frac{1}{2}}|exp\{-\frac{t^2}{8\sigma^2}\} \le 5^d exp\{-\frac{t^2}{8\sigma^2}\}$$

The desired bound is obtained by setting the right hand side equal to  $\sigma \in (0, 1)$  and solve for tNote: we have already seen from HW1 that under some regularity condition [Y10]:

$$|\hat{\Sigma}_n - \Sigma||_{\infty} \le C\sqrt{\frac{t + \log\delta}{n}}$$

with probability at least  $1 - e^{-t}$ , where  $\hat{\Sigma}_n$  is the empirical covariance matrix.

### 8.2 Matrix norm

**Definition 8.4** (Operator Norm) Let  $A \in \mathbb{R}^{m \times n}$ ,  $rank(A) = r \leq \min\{m, n\}$ . The singular value decomposition (SVD) of A is  $A = UDV^T$  where

- 1.  $D = diag(\sigma_1, \dots, \sigma_r)$ .  $\sigma_1 \ge \dots \ge \sigma_r > 0$  are the singular values of A.
- 2.  $U \in \mathbb{R}^{m \times r}$  whose columns are orthonormal and are called singular vectors
- 3.  $V \in \mathbb{R}^{n \times r}$  whose columns are orthonormal and are called singular vectors

Then  $AA^Tu_j = \sigma_j^2 u_j$ ,  $A^TAv_j = \sigma_j^2 v_j$  where  $u_j, v_j$  are the *j*-th column of U and V respectively. The operator norm of A is:

$$||A||_{op} = \max_{i} \sigma_{i} = \max_{x \in \mathbb{R}^{n} \setminus \{0\}} \frac{||Ax||}{||X||} = \max_{\substack{x \in \mathbb{S}^{n-1} \\ y \in \mathbb{S}^{n-1}}} x^{T} A y$$

Remarks:

- When  $A \in S^n$ (symmetric),  $||A||_{op} = \max_{x \in \mathbb{S}^{n-1}} |x^T A x|$ .
- We say  $A \in S^n_+$  (positive semi-definite (PSD)) if and only if  $\forall x \in \mathbb{R}^n, x^T A x \ge 0$ . As an example, any covariance matrix  $\Sigma$  is PSD because  $\mathbb{V}[a^T x] = a^T \ge 0, \forall a \in \mathbb{R}^n$
- If  $A \in S^n_+$ ,  $\sigma_i = \lambda_i$  where  $\lambda_i$ 's are the eigenvalues of A,  $||A||_{op} = \max_i \lambda_i = \max_{x \in \mathbb{S}^{n-1}} x^T A x$

The following two types of norms are also common in practice.

Definition 8.5 (Frobenius Norm)

$$||A||_F = \sqrt{\sum_{i,j=1} A_{ij}^2}$$

Definition 8.6 (p-Schatten Norm)

$$||A||_p = (\sum_{i=1}^{n \wedge m} \sigma_i^p(A))^{1/p}$$

where  $\sigma_i(A)$ 's are the singular values of A. When p = 1,  $||A||_p$  is the nuclear norm. When  $p = \infty$ ,  $||A||_p$  is the spectral norm.

The following two inequality are often useful in practice:

#### Lemma 8.7

$$||Ax|| \le ||A||_{op}||x|| \ \forall x$$

**Lemma 8.8** (Weyl's inequality) Assume  $A, B \in \mathbb{R}^m$  have singular values  $\sigma_i(A), \sigma_j(B)$  for  $i = 1, \dots, n \land m$ ;  $j = 1, \dots, n \land m$ , then:

$$\max_{i} |\sigma_i(A) - \sigma_i(B)| \le ||A - B||_{op}$$

**Corollary**  $||A - B||_{op} \to 0 \Rightarrow |x^T(A - B)y| \to 0$  uniformly for every  $x \in \mathbb{S}^{m-1}, y \in \mathbb{S}^{n-1}$ .

## 8.3 Covariance matrix estimation in the operator norm

**Theorem 8.9** Let  $X_1, \dots, X_n$  be iid samples from a distribution with mean 0 and covariance matrix  $\Sigma$ . Assume  $X_i \in SG_d(\sigma^2)$  and are centered. Let  $\hat{\Sigma}_n = \frac{1}{n} \sum_{i=1}^n X_i X_i^T$ . Then there exists a universal constant C > 0 s.t.

$$\mathbb{P}(\frac{||\hat{\Sigma}_n - \Sigma||_{op}}{\sigma^2} \ge C \max\{\sqrt{\frac{d + \log(\frac{2}{\delta})}{n}}, \frac{d + \log(\frac{2}{\delta})}{n}\}) \le \delta, \ \delta \in (0, 1)$$

**Remark**: Theorem 8.9 indicates that  $\hat{\Sigma}_n \xrightarrow{p} \Sigma$  with respect to the operator norm requires  $\frac{d}{n} \to 0$ **Proof ideas**: Use discretization and sub exponential concentration bound. Recall that  $X \in SG(\sigma^2) \Rightarrow X^2 - \mathbb{E}[X^2] \in SE(16^2\sigma^4, 16\sigma^2).$ 

**Lemma 8.10** Let  $A := \hat{\Sigma}_n - \Sigma \in S^n$  and  $N_{\epsilon}$  be the  $\epsilon$ -net of  $\mathbb{S}^{d-1}$  for  $\epsilon \in (0, \frac{1}{2})$ , then:

$$||A||_{op} = \max_{x \in \mathbb{S}^{n-1}} |x^T A x| \le \frac{1}{1 - 2\epsilon} \max_{y \in N_{\epsilon}} |y^T A y|$$

# References

- [VK14] V. KOLTCHINSKII and K. LOUNICI, "CONCENTRATION INEQUALITIES AND MOMENT BOUNDS FOR SAMPLE COVARIANCE OPERATORS," *arXiv preprint*, 2014, ArXiv:1405.2468.
- [Y10] Y. MING, "HIGH DIMENSIONAL INVERSE COVARIANCE MATRIX ESTIMATION VIA LINEAR PRO-GRAMMING," Journal of Machine Learning Research, 2010, pp. 2261–2286.