## 36-710: Advanced Statistical Theory

## Lecture 17: March 26

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This lecture covers slow and fast rates for the Lasso with the restricted eigenvalue condition.

## 17.1 Slow rate for the Lasso

Last lecture we proved Theorem 17.1 which is known as slow rate for the Lasso.

**Theorem 17.1** (Theorem 7.5 in the book.). Let  $\mathbf{X} \in \mathbb{R}^{nxd}$  and  $\beta^* \in \mathbb{R}^d$  an unknown vector in the standard linear model  $y = \mathbf{X}\beta^* + \epsilon$ . If  $\lambda = \lambda_n \geq \frac{\|\mathbf{X}^T \epsilon\|_{\infty}}{n}$ , then any lasso solution  $\hat{\beta}_{lasso}$  satisfies,

$$MSE(\hat{\beta}_{lasso}) = \frac{\left\| \mathbf{X}(\hat{\beta}_{lasso} - \beta^*) \right\|^2}{n} \le 4 \left\| \beta^* \right\|_1 \lambda_n.$$

Let's compare  $MSE(\hat{\beta}_{lasso})$  to the  $MSE(\hat{\beta}_{bss})$  for best subset selection.

$$MSE(\hat{\beta}_{bss}) \lesssim \|\beta^*\|_0 \sigma^2 \frac{\log(ed/\delta)}{n}$$
 with probability  $\geq 1 - \delta$ 

 $MSE(\hat{\beta}_{BSS})$  is almost optimal except the log term. Note that Theorem 17.1 is a deterministic result. In order to turn it into a more practical one, we ask the question: When can we get w.h.p. that

$$\lambda_n \ge \frac{\left\|\mathbf{X}^T \boldsymbol{\epsilon}\right\|_{\infty}}{n} = \max_{i=1,\dots,n} \frac{|X_i^T \boldsymbol{\epsilon}|}{n}$$

where  $X_i$  is the *i*'th column of **X**? Assume  $\max_i ||X_i|| \leq \sqrt{Cn}$  for some C > 0 (\*), then,

$$P\left(\frac{\|\mathbf{X}^{T}\boldsymbol{\epsilon}\|_{\infty}}{n}\geq\right) = P\left(\max_{i=1,\dots,n}\frac{|X_{i}^{T}\boldsymbol{\epsilon}|}{n}\geq t\right)$$

$$= P\left(\max_{i=1,\dots,n}|X_{i}^{T}\boldsymbol{\epsilon}|\geq tn\right)$$

$$\leq \sum_{i=1}^{d} P\left(|X_{i}^{T}\boldsymbol{\epsilon}|\geq tn\right)$$

$$= \sum_{i=1}^{d} P\left(\frac{|X_{i}^{T}\boldsymbol{\epsilon}|}{\|X_{i}\|}\geq\frac{tn}{\|X_{i}\|}\right)$$

$$\left(\frac{|X_{i}^{T}\boldsymbol{\epsilon}|}{\|X_{i}\|}\in SG(\sigma^{2}) \text{ is unit vector}\right)$$

$$\leq 2dexp\left(-\frac{t^{2}n}{2\sigma^{2}C}\right)$$
(Using bound on SG variable and (\*))

Then set RHS equal to  $\delta$  and solve for t. Setting  $t = \lambda_n = \sqrt{\frac{2\sigma^2 C}{n}(\log(2d) + \log(1/\delta))}$  yields that  $\lambda_n \geq \frac{\|\mathbf{x}^T \boldsymbol{\epsilon}\|_{\infty}}{n}$  with probability  $\geq 1 - \delta$ . Plug this into MSE for Lasso, then with probability  $\geq 1 - \delta$ ,

$$MSE(\hat{\beta}_{lasso}) \le 4 \left\|\beta^*\right\|_1 \sqrt{\frac{2\sigma^2 C}{n} (\log(2d) + \log(1/\delta))}.$$

Note that this MSE still vanishes, but at a slower rate than BSS. It's off by a square-root. To further improve on  $MSE(\hat{\beta}_{lasso})$ , we put additional assumptions on the design matrix.

## 17.2 Fast rates for the Lasso

To obtain a faster rate of convergence we need stronger assumptions on  $\frac{\mathbf{X}^T X}{n}$ .

**Definition 17.2** (Restricted Eigenvalue Condition  $(Re(\alpha, \kappa, S))$ ). For  $S \subseteq \{1, \ldots, d\}$ ,  $S \neq \emptyset$  and  $\alpha \ge 1$ , let

$$C_{\alpha}(S) = \{ \Delta \in \mathbb{R}^d : \|\Delta_{S^c}\|_1 \le \alpha \|\Delta_S\|_1 \} \quad where \quad S^c = \{1, \dots, d\} \setminus S.$$

Then, we say that  $X \in \mathbb{R}^{n \times d}$  satisfies REC with respect to  $S \subseteq \{1, \ldots, d\}$  and parameters  $\alpha \geq 1$ and  $\kappa > 0$  if

$$\frac{1}{n} \left\| \mathbf{X} \Delta \right\|^2 \ge \kappa \left\| \Delta \right\|^2 \qquad \forall \Delta \in C_{\alpha}(S)$$

**Intuition:** Set  $\Delta = \hat{\beta}_{lasso} - \beta^*$ . Then  $\frac{1}{n} \|\mathbf{X}\Delta\|^2$  is the MSE of  $\hat{\beta}_{lasso}$ . The function  $z \to \frac{\|\mathbf{X}z\|^2}{n}$  can be very flat in the sense that  $\frac{\|\mathbf{X}\Delta\|^2}{n}$  can be small but  $\Delta$  can still be large.

If  $\frac{\mathbf{X}^T \mathbf{X}}{n}$  has a minimum eigenvalue bounded away from zero, then

$$\left\|\Delta\right\|^2 \leq \frac{\frac{\|\mathbf{X}\Delta\|^2}{n}}{\lambda_{min}\left(\frac{\mathbf{X}^T\mathbf{X}}{n}\right)}$$

which requires that  $\frac{\Delta^T \mathbf{X}^T \mathbf{X} \Delta}{n} \ge \lambda_{min} > 0, \ \forall \Delta \in \mathbb{R}^d.$ 

**Theorem 17.3** (Theorem 7.2 in the book.). Assume that X satisfies REC w.r.t.  $Re(\alpha, \kappa, S)$  where S is the support of  $\beta^*$ , then if  $\lambda_n \geq 2 \frac{\|\mathbf{x}^T \boldsymbol{\epsilon}\|_{\infty}}{n}$ , any lasso solution satisfies,

$$MSE(\hat{\beta}_{lasso}) = \frac{\left\|\mathbf{X}(\hat{\beta}_{lasso} - \beta^*)\right\|^2}{n} \le 9\lambda_n^2 \frac{|S|}{\kappa}$$

where  $9\lambda_n^2 \frac{|S|}{\kappa} \simeq \frac{\|\beta^*\|_0}{\kappa} \frac{\sigma^2(\log(d) + \log(1/\delta))}{n}$  and

$$\left\|\hat{\beta}_{lasso} - \beta^*\right\| \le \frac{3}{\kappa} \sqrt{|S|} \lambda_n$$

Also note that  $|S| = \|\beta^*\|_0$ .

*Proof.* We first need to show that  $\Delta = \hat{\beta} - \beta^* \in C_3(S)$ . By optimality of  $\hat{\beta}$ , it holds that

$$\frac{1}{2n} \left\| Y - \mathbf{X} \hat{\beta} \right\|^2 + \lambda_n \left\| \hat{\beta} \right\|_1 \le \frac{1}{2n} \left\| Y - \mathbf{X} \beta^* \right\|^2 + \lambda_n \left\| \beta^* \right\|_1.$$

By rearraging, we obtain that,

$$\frac{1}{2n} \left\| \mathbf{X} \Delta \right\|^2 \le \frac{\epsilon^T \mathbf{X} \Delta}{n} + \lambda_n (\left\| \beta^* \right\|_1 - \left\| \hat{\beta} \right\|_1).$$

Since  $\beta^*$  is S-sparse, i.e.  $supp(\beta^*) = S$ ,

$$\begin{split} ||\beta^*||_1 - ||\hat{\beta}||_1 &= ||\beta^*_S||_1 - ||\beta^*_S + \Delta_S||_1 - ||\hat{\beta}_{S^c}||_1 \\ &= ||\beta^*_S||_1 - ||\beta^*_S + \Delta_S||_1 - ||\Delta_{S^c}||_1 \\ (\text{Since } ||\hat{\beta}||_1 &= ||\hat{\beta}_S||_1 + ||\hat{\beta}_{S^c}||_1) \end{split}$$

So, by Hölder's Inequality, it holds that,

$$\frac{1}{n} \left\| \mathbf{X} \Delta \right\|^2 \le 2 \frac{\left\| \mathbf{X}^t \epsilon \right\|_{\infty}}{n} \left\| \Delta \right\|_1 + 2\lambda_n$$

Now, using the facts that

1.  $\|\beta_S^* + \Delta_S\|_1 \ge \|\Delta_S\|_1 - \|\Delta_S\|_1$ 2.  $2\frac{\|\mathbf{x}^T \boldsymbol{\epsilon}\|_{\infty}}{n} \le \lambda_n.$ 

$$\frac{1}{n} \|\mathbf{X}\Delta\|^2 \le [||\Delta_S||_1 + ||\Delta_{S^c}||_1 + 2||\Delta_S||_1 - 2||\Delta_{S^c}||_1] = \underbrace{\lambda_n}_{\Delta \in C_3(S)} \underbrace{(3||\Delta_S||_1 - ||\Delta_{S^c}||_1)}_{\ge 0}$$

Next,

$$\begin{split} \lambda_n \left( 3 \| \Delta_S \|_1 - \| \Delta_{S^c} \|_1 \right) &\leq \lambda_n 3 \| \Delta_S \|_1 \\ &\leq 3\lambda_n \sqrt{|S|} \| \Delta_S \| \\ &\leq 3\lambda \sqrt{|S|} \frac{\| \mathbf{X} \Delta_S \|}{\sqrt{n\kappa}} \end{split}$$
(by REC)

Thus,

$$\frac{1}{\sqrt{n}} \|\mathbf{X}\Delta\| \le 3\lambda_n \sqrt{\frac{|S|}{\kappa}}$$

taking square of both sides gives

$$\frac{1}{n} \left\| \mathbf{X} \Delta \right\|^2 \le 9\lambda_n^2 \frac{|S|}{\kappa}$$

as claimed. Similarly, for the second part

$$\sqrt{\kappa} \|\Delta\| \underset{REC}{\leq} \frac{\|\mathbf{X}\Delta\|}{\sqrt{n}} \underset{\text{Above bound}}{\leq} 3\lambda_n \frac{\sqrt{|S|}}{\sqrt{\kappa}}$$

So, with probability  $\geq 1 - \delta$ ,

$$||\Delta|| = ||\hat{\beta} - \beta^*|| \le 3\lambda_n \frac{\sqrt{|S|}}{\kappa} \lesssim \sqrt{\|\beta^*\|_{\infty}} \sigma \sqrt{\frac{(\log(d) + \log(1/\delta)}{\kappa}}$$