36-755: Advanced Statistical Theory 1

Lecture 11: October 10

Lecturer: Alessandro Rinaldo

Scribes: Pengtao Xie

Fall 2016

Note: LaTeX template courtesy of UC Berkeley EECS dept.

Disclaimer: These notes have not been subjected to the usual scrutiny reserved for formal publications. They may be distributed outside this class only with the permission of the Instructor.

This lecture's notes illustrate some uses of various IATEX macros. Take a look at this and imitate.

11.1 Persistence

Setup: Z_1, \dots, Z_n are i.i.d samples drawn from distribution P, where $Z_i = (Y_i, X_i)$ with $Y_i \in \mathbb{R}$ and $X_i \in \mathbb{R}^d$, and $Y_i = f(X_i) + \epsilon$. f can be any function. We want to predict Y using vector X. We are only using linear predictors. Formally, for any $\beta \in \mathbb{R}^d$, let

$$R_P(\beta) = \mathbb{E}_P[(Y - X^\top \beta)^T]$$
(11.1)

We assume $\operatorname{cov}[X] = \Sigma$ is non-singular, then the problem

$$\min_{\beta \in \mathbb{R}^d} R_P(\beta) \tag{11.2}$$

has unique solution $\beta^* = \Sigma^{-1} \alpha$ where $\alpha = \mathbb{E}[YX]$.

Suppose we have a sequence $\{P_n\}$ of probability distribution for $Z = (Y, X) \in \mathbb{R}^{d+1}$ where d = d(n). We also have a sequence of sets $\{B_n\}$ where $B_n \subset \mathbb{R}^{d(n)}$. For each n, let the optimal constrained parameters be

$$\beta_n^* \in \operatorname{argmin}_{\beta \in B_n} R_{P_n}(\beta) \tag{11.3}$$

Example of B_n : (1) $B_n = \{\theta \in \mathbb{R}^{d(n)}, \|\beta\|_1 \le b_n\}$ where $b_n > 0$; (2) $B_n = \{\theta \in \mathbb{R}^{d(n)}, \|\beta\|_0 \le k_n\}$

Definition of persistence: given a sequence $\{(P_n, \beta_n^*)\}$, a sequence of estimators $\{\hat{\beta}_n\}$ is persistent if $R_{P_n}(\hat{\beta}_n)$ converges to $R_{P_n}(\beta_n^*)$ in probability.

We will be looking at

$$\hat{\beta}_n = \operatorname{argmin}_{\beta \in B_n} \hat{R}(\beta) \tag{11.4}$$

where

$$\hat{R}(\beta) = \frac{1}{n} \sum_{i=1}^{n} (y_i - x_i^{\top} \beta)$$
(11.5)

Let $\widetilde{\Sigma} = \operatorname{cov}[Z]$ and \hat{Z} be the empirical covariance. Assume that $\|\widetilde{\Sigma} - \hat{\Sigma}\|_{\infty} = \max_{ij} |\widetilde{\Sigma}_{ij} - Si\hat{g}ma_{ij}| \leq \Delta_n(\delta)$ with probability $1 - \delta$ for all n and P_n . For $\beta \in \mathbb{R}^{d+1}$, let $\tilde{\beta} = (1, -\beta) \in \mathbb{R}^{d+1}$, then $y - x^{\top}\beta = z^{\top}\tilde{\beta}$. Let $\tilde{B}_n = \{(-1, \beta) \in \mathbb{R}^{d+1}, \beta \in B_n\}$, then $R_P(\beta) = R_p(\tilde{\beta})$.

Theorem: Assume $d = n^{\alpha}$ where $\alpha > 0$. Then

$$R_{P_n}(\hat{\beta}) \le R_{P_n}(\tilde{\beta}^*) + 2\Delta_n (b_n + 1)^2$$
 (11.6)

Proof: $R_{P_n}(\tilde{\beta}) = \tilde{\beta}^\top \tilde{\Sigma} \tilde{\beta}$ and $\hat{R}_{\tilde{\beta}} = \tilde{\beta}^\top \hat{\Sigma} \tilde{\beta}$. Then $\forall \tilde{\beta} \in \mathbb{R}^{d+1}$ and P_n , we have

$$\begin{aligned} &|R_{P_n}(\tilde{\beta}) - \hat{R}_{\tilde{\beta}}| = |\tilde{\beta}^\top (\tilde{\Sigma} - \hat{\Sigma})\tilde{\beta}| \\ &\leq \|\tilde{\Sigma} - \hat{\Sigma}\|_{\infty} \|\tilde{\beta}\|_1 \text{(Holder Inequality)} \\ &\leq \Delta_n(\delta)(b_n + 1)^2 \end{aligned} \tag{11.7}$$

Then

$$R_{P_n}(\tilde{\beta}_n) \le \hat{R}(\hat{\beta}_n) + \Delta_n(\delta)(b_n + 1)^2$$

$$\le \hat{R}(\hat{\beta}_n^*) + \Delta_n(\delta)(b_n + 1)^2$$

$$\le \hat{R}_{P_n}(\hat{\beta}_n^*) + \Delta_n(\delta)(b_n + 1)^2$$
(11.8)

Remark: If $\Delta_n(\delta) \preceq \sqrt{\frac{\log d}{n} + \frac{\log(1/\delta)}{n}} \preceq \sqrt{\frac{\log n}{n}}$ If $d = n^{\alpha}$, $\delta = \frac{1}{n}$ and $\tilde{\beta}_n = \{\tilde{\beta} \in \mathbb{R}^{d+1} | \|\tilde{\beta}\|_1 \leq b_n + 1\}$. Then $\hat{\tilde{\beta}}_n$ is persistent if $b_n = o((\frac{n}{\log n})^{\frac{1}{4}})$

11.2 PCA

Let $X \in \mathbb{R}^d$ be a random vector with $cov[X] = \Sigma$. Let $\lambda_i(\Sigma)$ be the eigenvalue of Σ and u_i be the eigenvector associated with $\lambda_i(\Sigma)$. Assume $\lambda_{\max} = \lambda_1(\Sigma) \geq \cdots \geq \lambda_d(\Sigma) \geq 0$.

PCA has several interpretations.

Optimal Linear Subspace: what is the direction $v \in \mathbb{S}^{d-1}$ such that $\operatorname{var}[v^{\top}X]$ is maximal?

 $v^* = \operatorname{argmax}_{v \in \mathbb{S}^{(d-1)}} \operatorname{var}[v^\top X]$ is the eigenvector associated to $\lambda_{\max}(\Sigma)$.

More generally, let $V_{d \times r} = \{V_{d \times r} \text{ with orthogonal columns}\}$. The optimal solution of

$$\operatorname{argmax}_{V \in V_{d \times r}} \mathbb{E}[\|V^{\top}X\|^2] \tag{11.9}$$

is the first r eigenvectors.

Low-rank Approximation: We want to find matrix Z^* such that

$$Z^* \in \operatorname{argmin}_{F} \|\Sigma - Z\|_F^2$$

s.t. rank(Z) = r (11.10)

Then $Z^* = \sum_{i=1}^r \lambda_i \mu_i \mu_i^\top$ and $\|Z^* - \Sigma\|_F^2 = \sum_{j=i+1}^d \lambda_j^2$

Subspace: suppose we want to find subspace S of R^d of dimension $r \leq d$.

$$\mathbb{E}\|X - \Pi_S X\|^2 \tag{11.11}$$

where Π_S is the orthonormal projection of X onto S. Then $\Pi_S = V_r V_r^{\top}$ where the columns of V_r are r largest eigenvectors.

The challenges are that we need to estimate eigenvalues and eigenvectors well.