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Lecturer: Alessandro Rinaldo <arinaldo@cmu.edu>

Scribe: Neil Spencer

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9.1 Lasso

Consider a regression framework where Y is an $n \times 1$ vector, X is a $n \times d$ matrix, θ^* is a $d \times 1$ vector, and ϵ is a $n \times 1$ vector. Further assume that

$$Y = X\theta^* + \epsilon$$

with $\epsilon \in \text{SG}_n(\sigma^2)$. In the LASSO, we use estimate $\hat{\theta}$ to estimate θ , where $\hat{\theta}$ is the solution to

$$\hat{\theta} \in \operatorname{argmin}_{\theta \in \mathbb{R}^d} \left(\frac{1}{2n} \|Y - X\theta\|^2 + \lambda_n \|\theta\|_1 \right). \quad (9.1)$$

Equation (1) defines a convex optimization problem that produces sparse solutions depending on λ_n . The parameter λ_n is chosen by the user. It can be thought of as $\lambda(n, d, \sigma)$ because the choice will depend on those values.

Equation (1) has solutions for both $d \leq n$ and $d > n$. There can be multiple optimal solution $\hat{\theta}$, but the maximizing value $X\hat{\theta}$ is unique. For a discussion of the uniqueness of solutions to the Lasso problem, see [1].

The basic inequality [2] is a useful inequality for proving results pertaining to the Lasso. It is given below as Lemma 1.1. It is used to prove Theorem 1.2.

Lemma 9.1. *In the Lasso set-up, if θ^* is the true parameter value and $\hat{\theta}$ is the lasso solution, then*

$$\frac{1}{2n} \left\| X(\hat{\theta} - \theta^*) \right\|^2 \leq \epsilon^T \frac{X(\hat{\theta} - \theta^*)}{n} + \lambda_n (\|\theta^*\|_1 - \|\hat{\theta}\|_1).$$

Proof.

$$\begin{aligned} \frac{1}{2n} \left(\|\epsilon\|^2 + \left\| X(\hat{\theta} - \theta^*) \right\|^2 - 2\epsilon^T X(\hat{\theta} - \theta^*) \right) + \lambda_n \|\hat{\theta}\|_1 &= \frac{1}{2n} \left\| \epsilon - X(\hat{\theta} - \theta^*) \right\|^2 + \lambda_n \|\hat{\theta}\|_1 \\ &= \frac{1}{2n} \left\| Y - X\hat{\theta} \right\|^2 + \lambda_n \|\hat{\theta}\|_1 \\ &\leq \frac{1}{2n} \|Y - X\theta^*\|^2 + \lambda_n \|\theta^*\|_1 \\ &= \frac{1}{2n} \|\epsilon\|^2 + \lambda_n \|\theta^*\|_1 \end{aligned}$$

□

Theorem 9.2. *If*

$$\lambda_n \geq \left\| \frac{X^T \epsilon}{n} \right\|_{\infty} = \max_{J=1, \dots, d} \left| \frac{X_J^T \epsilon}{n} \right|,$$

then any Lasso solution satisfies

$$\left\| \frac{X(\hat{\theta} - \theta^*)}{n} \right\|^2 \leq 4 \|\theta^*\|_1 \lambda_n.$$

Proof. Lemma 1.1 provides

$$\begin{aligned} \frac{1}{2n} \left\| X(\hat{\theta} - \theta^*) \right\|^2 &\leq \epsilon^T \frac{X(\hat{\theta} - \theta^*)}{n} + \lambda_n (\|\theta^*\|_1 - \|\hat{\theta}\|_1) \text{ by Basic inequality (Lemma 1.1)} \\ &\leq \frac{1}{n} \|X^T \epsilon\|_{\infty} \|\hat{\theta} - \theta^*\|_1 + \lambda_n (\|\theta^*\|_1 - \|\hat{\theta}\|_1) \text{ by Holder Inequality} \\ &\leq \frac{1}{n} \|X^T \epsilon\|_{\infty} (\|\theta^*\|_1 + \|\hat{\theta}\|_1) + \lambda_n (\|\theta^*\|_1 - \|\hat{\theta}\|_1) \text{ by Triangle Inequality} \\ &\leq \left(\frac{1}{n} \|X^T \epsilon\|_{\infty} - \lambda_n \right) \|\hat{\theta}\|_1 + \left(\frac{1}{n} \|X^T \epsilon\|_{\infty} + \lambda_n \right) \|\theta^*\|_1 \\ &\leq 2\lambda_n \|\theta^*\|_1. \end{aligned}$$

□

What is a good choice for λ_n ?

Recall that $\epsilon \in \text{SG}(\sigma^2)$ and assume $\max_{J=1 \dots d} \|X_J\| \leq C\sqrt{n}$ for some $C > 0$. Then for $t > 0$,

$$\begin{aligned} \mathbb{P} \left(\left\| \frac{\epsilon^T X}{n} \right\|_{\infty} \geq t \right) &\leq \mathbb{P} (\max_J |X_J^T \epsilon| \geq tn) \\ &\leq \sum_J \mathbb{P} (|X_J^T \epsilon| \geq tn) \\ &\leq \sum_J \mathbb{P} \left(\frac{|X_J^T \epsilon|}{\|X_J\|} \geq \frac{tn}{\|X_J\|} \right) \\ &\leq 2d \exp \left(\frac{-t^2 n}{2\sigma^2 \max \|X_J\|^2} \right) \text{ by Subgaussianity} \\ &\leq 2d \exp \left(\frac{-t^2 n}{2\sigma^2 C^2} \right) \text{ because } \|X_J\|^2 < C^2 n \\ &\leq \delta \end{aligned}$$

if we choose

$$t = \lambda_n = \sqrt{\frac{2\sigma^2 C^2}{n} \left(\log \left(\frac{1}{\delta} \right) + \log(2d) \right)}.$$

Consider $\delta = 1/n$. Then with probability $1 - 1/n$,

$$\left\| \frac{X(\hat{\theta} - \theta^*)}{n} \right\|^2 \leq \|\theta^*\|_1 \sqrt{\frac{2\sigma^2 C (\log(n) + \log(2d))}{n}}$$

If $d \leq n$ and $\lambda_{\min}(X^T X/n) \geq C_{\min} > 0$. Then you can also get a bound for

$$\left\| \hat{\theta} - \theta^* \right\|^2 \leq \frac{\|\theta^*\|}{c_{\min} \lambda_n}$$

9.2 Getting Fast Rates for Lasso

In order to get “fast” rates for the lasso, there needs to be additional assumptions on X . These assumptions also provide consistency of estimation of θ .

A very useful condition is the *restricted eigenvalue condition*. In order to define the condition, we need to establish some notation. For $S \subset \{1, 2, \dots, d\}$ and $\alpha > 0$, define

$$C(\alpha, S) = \{ \Delta \in \mathbb{R}^d : \|\Delta_{S^c}\|_1 \leq \alpha \|\Delta_S\|_1 \}.$$

Definition 9.3. X satisfies the restricted eigenvalue (RE(α, κ)) condition over $S = \{1, \dots, d\} \neq \emptyset$ if

$$\frac{1}{n} \|X\Delta\|^2 \geq \kappa \|\Delta\|^2 \text{ for all } \Delta \in C(\alpha, S).$$

For intuition, think of Δ as $\hat{\theta} - \theta^*$. We want $\|X\Delta\|^2/n$ to be small. Note that if it is, this does necessarily mean that $\|\Delta\|^2$ is small. Especially if

$$\Delta \rightarrow \frac{\|X\Delta\|^2}{n} \tag{9.2}$$

is flat around $\hat{\theta} - \theta^*$. To prevent this, we need the function (9.2) to be very curved. This is true if

$$\frac{\|X\Delta\|^2}{n} \geq \kappa \|\Delta\|^2 \text{ for all } \Delta \in \mathbb{R}^d.$$

Unfortunately, this implies that $\lambda_{\min}(X^T X) \geq C_{\min} > 0$ if $d > n$, which is not possible. Instead, we consider the case where function (9.2) is curved only along certain directions.

These directions are $C(\alpha, S)$ where S is defined by the support of θ^* . That is, $s = \{j : \theta_j^* \neq 0\}$.

Theorem 9.4. Assume that

- the support of θ^* is S where $|S| = s > 0$.
- X satisfies RE($3, \kappa$) where $\kappa > 0$ with respect to S .
- $\lambda_n \geq 2 \|\epsilon^T X\|/n$.

Then any Lasso solution $\hat{\theta}$ satisfies

$$\frac{1}{n} \left\| X \left(\hat{\theta} - \theta^* \right) \right\|^2 \leq 9\lambda_n^2 s \kappa,$$

and

$$\left\| \hat{\theta} - \theta^* \right\| \leq \frac{3}{\kappa} \sqrt{s} \lambda_n.$$

Proof. First, we show that given our choice λ_n , $\hat{\Delta} = (\hat{\theta} - \theta^*) \in C(3, S)$. By the Basic inequality,

$$0 \leq \frac{1}{2n} \left\| X \hat{\Delta} \right\|^2 \leq \frac{\epsilon^T X \Delta}{n} + \lambda_n [\|\theta^*\|_1 - \|\hat{\theta}\|_1].$$

Since θ^* is S -sparse, we know

$$\begin{aligned} \|\theta^*\|_1 - \|\hat{\theta}\|_1 &= \|\theta_S^*\|_1 - \left\| \theta_S^* + \hat{\Delta}_S \right\|_1 - \left\| \hat{\theta}_{S^c} \right\|_1 \\ &= \|\theta_S^*\|_1 - \left\| \theta_S^* + \hat{\Delta}_S \right\|_1 - \left\| \hat{\Delta}_{S^c} \right\|_1. \end{aligned}$$

Plugging this into the Basic inequality yields:

$$\begin{aligned} 0 &\leq \frac{1}{2n} \left\| X \hat{\Delta} \right\|^2 \\ &\leq 2 \frac{\|X^T \epsilon\|_\infty}{n} \|\hat{\Delta}\|_1 + 2\lambda_n \left(\|\theta_S^*\|_1 - \|\theta_S^* + \hat{\Delta}_S\|_1 - \|\hat{\Delta}_{S^c}\|_1 \right) \\ &\leq 2 \frac{\|X^T \epsilon\|_\infty}{n} \|\hat{\Delta}\|_1 + 2\lambda_n \left(\|\hat{\Delta}_S\|_1 - \|\hat{\Delta}_{S^c}\|_1 \right) \text{ By triangle inequality} \\ &\leq \lambda_n \|\hat{\Delta}_S\|_1 + \lambda_n n \|\hat{\Delta}_{S^c}\|_1 + 2\lambda_n \|\hat{\Delta}_S\|_1 - 2\lambda_n \|\hat{\Delta}_{S^c}\|_1 \\ &= \lambda_n \left(3\|\hat{\Delta}_S\|_1 - \|\hat{\Delta}_{S^c}\|_1 \right) \\ &\Rightarrow \hat{\Delta} \in C(3, S). \end{aligned}$$

Note that the fourth line used the fact

$$\lambda_n \geq \frac{2 \|X^T \epsilon\|_\infty}{n}.$$

□

References

- [1] Ryan J Tibshirani. The lasso problem and uniqueness. *Electronic Journal of Statistics*, 7:1456–1490, 2013.
- [2] Sara A Van de Geer. *Applications of empirical process theory*, volume 91. Cambridge University Press Cambridge, 2000.